

## **Gender difference in the association between subjective socioeconomic mobility across life course and mortality at older ages: Results from the JAGES longitudinal study**

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1 **ABSTRACT**

2 **Background:** Socioeconomic mobility affects health throughout the life course. However, it is not  
3 known whether there are gender differences in the association between life-course subjective  
4 socioeconomic status (SSS) mobility and mortality at older ages.

5 **Methods:** Participants were 16,690 community-dwelling adults aged 65–100 years in the Japan  
6 Gerontological Evaluation Study (JAGES). Baseline information including demographic  
7 characteristics, depression and lifestyle factors were collected in 2010. Participants' vital status was  
8 confirmed in 2013 via linkage to death records. We categorized life-course socioeconomic mobility  
9 into the following categories: 'persistently high', 'downward mobility', 'upward mobility' and  
10 'persistently low'. Cox proportional hazard modeling was used to estimate hazard ratios (HR) for all-  
11 cause mortality.

12 **Results:** Mortality HRs for the 'downward' group were 1.37 (95% confidence interval [CI] 1.08–  
13 1.74) among men and 1.27 (95% CI 0.94–1.71) among women in comparison with the 'persistently  
14 high' group. Compared to the 'persistently low' group, the HRs for the 'upward' group were 0.54  
15 (95% CI 0.35–0.83) among women and 0.91 (95CI 0.73–1.24) among men. Associations were not  
16 changed after adjusting for objective socioeconomic status, but attenuated by depression.

17 **Conclusions:** 'Downward' mobility was associated with mortality among men, but not among  
18 women. Depression appeared to mediate the association. A protective effect of upward mobility was  
19 observed among women but not among men.

20

21 **Keywords:** Subjective socioeconomic status; trajectory; all-cause mortality; gender roles; Japan

22 **Running title:** Gender roles associated with SSS mobility and mortality

23

## 24 INTRODUCTION

25 Subjective socioeconomic status (SSS), defined as “a person’s belief about his location in a status  
26 order”<sup>1</sup>, has been shown to be correlated with health and illness independently from objective  
27 socioeconomic status (SES, including educational attainment, income and occupation)<sup>2</sup>. Singh-  
28 Manoux A et al. concluded that SSS can capture a perceived average of objective SES<sup>3</sup>. More recently,  
29 the biological pathways linking SSS to health have begun to be documented. People’s perceptions of  
30 their social status predict health behaviors, such as smoking<sup>4</sup> and sedentarism<sup>2</sup>, as well as some  
31 biomarkers, such as heart rate, sleep latency, cortisol habituation to stress, body fat distribution<sup>5</sup> and  
32 high waist circumference<sup>6</sup>. Additionally, SSS has been shown to be inversely related to metabolic  
33 syndrome<sup>6</sup>, insulin resistance<sup>7</sup>, coronary artery diseases, hypertension, diabetes and dyslipidemia<sup>8</sup>.  
34 As for older people, Instrumental Activities of Daily Living difficulties are also inversely related to  
35 SSS<sup>2</sup>. Presumably, through these pathways, low SSS is associated with excess risk of mortality<sup>9</sup>.  
36 The statistically significant associations between SSS and all outcomes above—except coronary  
37 artery disease, hypertension and diabetes—were observed even after adjusting for objective  
38 socioeconomic status.

39 Recent studies have begun to focus on the trajectories of SSS across the life course. For  
40 example, Burazeri et al.<sup>10</sup> found that downward mobility in SSS (i.e. high SSS in childhood/low SSS  
41 during mid-life) was associated with increased risk of acute coronary syndrome (ACS) for both men  
42 and women at around 60 years of age. In this study, financial loss appeared to mediate downward  
43 mobility and ACS. Depressive symptoms were not considered in the study, although depression might  
44 be a robust mediator of the association between SSS and ACS. Another study of Latino and Asian  
45 immigrants in the United States (US) found that downward subjective social mobility was associated  
46 with elevated risk of major depression<sup>11,12</sup>. However, to the best of our knowledge, no studies have  
47 examined the association between mobility in SSS and mortality among older people prospectively,  
48 independent from objective SES, and the mediating effect of depression in this association.

49 It is also not known whether there are gender differences in the association between mobility

50 in SSS and health outcomes. Previous studies on “objective” social class mobility have found gender  
51 differences. For example, Cambois et al. examined occupational class mobility and standardized  
52 mortality rates (SMR) stratified by sex in a French population aged 30–84 years old. In this study,  
53 occupational class was classified into six groups: ‘upper classes’, ‘craft, trade, etc.’, ‘farmers’,  
54 ‘clerks’, ‘manual workers’ and ‘inactive’. Among both men and women, individuals who were  
55 downwardly mobile had statistically significant higher mortality compared to ‘non-movers’. At the  
56 same time, men who moved upwards from ‘inactive’ to a higher occupational position also  
57 experienced higher mortality risk compared to non-movers. By contrast, among women, moving up  
58 from the ‘inactive’ group showed a protective effect on mortality risk<sup>13</sup>. A study in Poland reported  
59 that upward intergenerational educational mobility (comparing individuals to their fathers) was  
60 protectively associated with cardiovascular risk—as measured by the Framingham risk score (FRS)—  
61 only among men<sup>14</sup>. No study has investigated whether there are gender differences in the association  
62 between “subjective” social class mobility and mortality.

63 Gender roles closely mirror the cultural context. For example, in Japan’s traditional patriarchal  
64 culture, women can be college-educated, but still end up as home-makers and therefore do not benefit  
65 from upward educational mobility to the same extent as men<sup>15</sup>. We derive two theoretical predictions  
66 for the gender difference we should expect to observe in the association between mobility in SSS and  
67 mortality. In a rigidly traditional society (men are expected to be bread-winners and women are  
68 expected to stay at home), (i) upward social mobility may not be beneficial to women; and (ii) in the  
69 same society, downward mobility could be more “toxic” for men because the male is the breadwinner  
70 and the household’s fortunes are more closely linked to the male’s SES trajectory. In this present study,  
71 these hypotheses were examined.

72 The Japan Gerontological Evaluation Study (JAGES) is a large-scale, population-based cohort  
73 study established in 2010 that collected data on both childhood and current SSS among healthy older  
74 (65 years or older) community residents in Japan. Vital status was ascertained in the cohort via linkage  
75 to death registry data. Previously, using prospective JAGES data from 2003, we reported that lower

76 objective SES (educational attainment and household income) was associated with higher mortality  
77 <sup>16</sup>. However, the impact of differential trajectories of SSS on mortality risk has not been investigated  
78 in older Japanese adults. In the present study, our objectives were to examine whether trajectories of  
79 SSS across the life course, that is, subjective social status mobility from childhood to old age, are  
80 associated with the risk of mortality and whether there is a gender-related difference in the association.

81

## 82 **METHODS**

### 83 **Study participants**

84 Data from an ongoing prospective longitudinal cohort study, the Japan Gerontological Evaluation  
85 Study (JAGES), was analyzed. The cohort was established with the aim of elucidating the determinants  
86 of healthy ageing among community-dwelling seniors. The baseline survey was administered between  
87 August 2010 and January 2012, when survey questionnaires were mailed to 131,468 community-  
88 dwelling individuals aged  $\geq 65$  years who were physically and cognitively independent. People from  
89 24 municipalities throughout Japan were invited to participate. Participants were randomly selected  
90 from the official residence registries in 13 large municipalities; and in the remaining 11 smaller  
91 municipalities, census data of all eligible residents was sought. A total of 86,063 subjects returned the  
92 questionnaire (response rate: 65.5%).

93 Approximately one-fifth of the total sample (N=19,528) was randomly selected to receive a  
94 supplementary survey module inquiring about their childhood and current SSS. After excluding those  
95 who had missing SSS responses (N=2,832) as well as individuals who could not be linked to mortality  
96 register data (n=6), our analytical sample consisted of 16,690 subjects (7,944 men and 8,746 women).  
97 The JAGES protocol was approved by the Ethics Committee in Research of Human Subjects at Nihon  
98 Fukushi University (No. 10-05) and the Ethics Committee of Chiba University, Faculty of Medicine  
99 (No. 1777).

100

### 101 **Ascertainment of death**

102 Participants were matched to death records from the public long-term care insurance system, which  
103 is maintained by local municipal authorities throughout Japan. All deaths were reported and submitted  
104 by physicians to municipalities. In the analysis sample from 2013, 780 deaths were found, which was  
105 4.7% of total participants at baseline. Of these deaths, 524 were men (6.6% of 7,944 male participants),  
106 and 256 were women (2.9% of 8,746 female participants).

107

## 108 **Childhood and current SSS**

109 Childhood SSS was asked using the following question: “How would you rate your social status when  
110 you were aged 15 years compared to others in society at that time?”<sup>17</sup>. Possible responses were  
111 arranged on a 5-point Likert scale: ‘high’, ‘middle-high’, ‘middle’, ‘middle-low’, or ‘low’.

112 Current SSS was assessed by the question: “How would you rate your current social status in  
113 comparison with others in society?”, with the same 5-point Likert scale of responses.

114 Since a previous study showed that even subtle downward/upward mobility can be  
115 associated with health outcomes<sup>13</sup>, we subtracted the 5-point Likert scale response for childhood SSS  
116 from the same 5-point Likert scale response for adult SSS, and categorized the resulting numbers into  
117 ‘downward SSS trajectory’ if the calculated result was -4 to -1, ‘no change’ if 0 and ‘upward SSS  
118 trajectory’ if +1 to +4. Subsequently, ‘no change’ was dichotomized into two groups and the  
119 individual was defined as ‘persistently high’ if both childhood and current SSS were ‘high’, ‘middle-  
120 high’ or ‘middle’, and ‘persistently low’ if both SSS were ‘middle-low’ or ‘low’.

121

## 122 **Covariates**

123 Self-reported adult height, current income, education and marital status were included as potential  
124 confounding factors. Additionally, depression and lifestyle factors were included as potential  
125 mediators of the association between SSS mobility and mortality. Adult height was used as a proxy  
126 of childhood adversity such as childhood nutrition and disease status<sup>18,19</sup>. Adult height was  
127 categorized into five groups in 5cm intervals specific to each gender: for men: <155, 155–159.9, 160–

128 164.9, 165–169.9, and  $\geq 170$ cm; and for women: <145, 145–149.9, 150–154.9, 155–159.9, and  
129  $\geq 160$ cm<sup>17</sup>. Education was assessed as years of schooling (0–9, 10–12, or  $\geq 13$  years) and annual  
130 household income was equalized by dividing the gross income by the square root of the number of  
131 household members (< 2.00, 2.00–3.99, or  $\geq 4.00$  million yen). We decided not to include  
132 occupational class as a covariate after performing the sensitivity analysis. Marital status was  
133 categorized as married, widowed, divorced, unmarried or others. The 15-item short form of the  
134 Geriatric Depression Scale (GDS) (Japanese version) was used to assess depression and categorized  
135 as no depression (0–4), moderate depression (5–9), and depression (10+) <sup>20,21</sup>. Lifestyle factors  
136 included smoking status (current smoker/ex-smoker or non-smoker), alcohol intake (current  
137 drinker/ex-drinker or non-drinker), walking time per day (<30min, 30 $\leq$ ) and body mass index (BMI)  
138 (underweight (<18.5), normal (18.5–24.9), overweight (25.0–29.9) or obesity (30.0 $\leq$ )) <sup>22</sup>.

139

#### 140 **Statistical analysis**

141 Data was analyzed by gender because self-rated health <sup>23</sup> and depressive symptoms <sup>24,25</sup> differed  
142 between men and women in our sample <sup>26,27</sup>. Cox proportional hazards models were used to estimate  
143 hazard ratios (HR) and 95% confidence intervals (CI) for all-cause mortality during the 3-year follow-  
144 up period. The following sequential multivariable-adjusted models were constructed: model 1 was  
145 adjusted for age, model 2 was additionally adjusted for adult height, adult objective SES (educational  
146 and income level) and marital status, model 3 was additionally adjusted for depression, and model 4  
147 was additionally adjusted for lifestyle factors. Analysis was performed using Stata/SE version 13.1  
148 (Stata Corp, College Station, TX 77845, USA).

149

#### 150 **RESULTS**

151 The overall and gender-stratified characteristics of the study subjects are described in Table 1. Among  
152 all participants, 25.7% of men and 31.5% of women reported both a high childhood SSS and high  
153 current SSS, while 18.7% of men and 14.3% of women reported low current SSS and childhood SSS.

154 Twenty percent of men and 30.0% of women were categorized as ‘downward mobility’ while 35.6%  
155 of men and 24.2% of women were categorized as ‘upward mobility’ (Table 1).

156 Among men, after adjustment for age, the HR from the ‘downward mobility’ group was 1.37  
157 (95% CI: 1.08–1.74) compared to those in the ‘persistently high’ group (Model 1, Table 2). Even after  
158 adjustment for adult height, adult objective SES and marital status, the HR of ‘downward mobility’  
159 was significantly higher compared to ‘persistently high’ (HR 1.35; 95% CI 1.05–1.72, Model 2).  
160 Additional adjustment for depression attenuated the association between ‘downward mobility’ and  
161 mortality, which became non-significant (HR 1.22; 95% CI 0.95–1.56, Model 3). Additional  
162 adjustment for lifestyle factors showed a non-significant association (HR 1.22; 95% CI 0.95–1.56,  
163 Model 4). On the other hand, compared to the ‘persistently low’ group, no significant association was  
164 found between ‘upward mobility’ and mortality among men.

165 Interestingly, the trend was opposite in women (Table 3). Women from the downward  
166 mobility group showed no significant excess mortality risk compared to the ‘persistently high’ group  
167 either before (HR 1.27; 95% CI 0.94–1.71, Model 1) or after adjustment for age, adult height,  
168 objective SES and other covariates (HR 1.08; 95%CI 0.79–1.47, Model 4). Although the result was  
169 not statistically significant, the point estimate of the HR showed that adjustment for depression  
170 considerably attenuated the effect of downward mobility among women also. We additionally  
171 conducted a check for the interaction between gender by mobility in the SSS and mortality. The result  
172 confirmed that downward mobility was harmful only among men. However, compared to  
173 ‘persistently low’, the age-adjusted HR of the ‘upward mobility’ group was 0.54 (95% CI: 0.35–0.83,  
174 Model 5). After adjustment for adult height, adult objective SES and marital status, the HR for  
175 ‘upward mobility’ was slightly attenuated to 0.58 (95% CI: 0.38–0.90, Model 6). In the final model,  
176 after additional adjustment for depression and other covariates, the HR for ‘upward mobility’ was  
177 attenuated to statistical non-significance (HR 0.68; 95% CI: 0.44–1.07, Model 8).

178

179 **DISCUSSION**



180 In this analysis of life-course trajectories of SSS among Japanese older adults, we found four notable  
181 patterns: (a) ‘downward mobility’ increased the risk of mortality compared to the ‘persistently high’  
182 group only among men; (b) ‘upward mobility’ decreased the risk of mortality only among women,  
183 with no effect seen among men; (c) depression appeared to mediate the association between mobility  
184 in SSS and mortality, and finally, (d) objective SES in later life did not fully explain the relationship  
185 between mobility in SSS and mortality.

186 Our finding of excess mortality among downwardly mobile groups is consistent with  
187 previous studies in other populations and other age groups<sup>10-12</sup>. However, it was observed only among  
188 men in the older Japanese population. Downward mobility increases the risk of poor health for men  
189 because of a subjective sense of deprivation in adult life relative to their childhood status. This might  
190 be less salient for women in Japanese society because their economic fortunes are more tightly linked  
191 to their spouse’s status. As shown in Table 1, over 90% of women in the sample were married or  
192 previously married (now widowed). Also, as mentioned above, the gender difference here is probably  
193 real because the formal test of interaction was statistically significant only among men. In addition to  
194 that, depression is probably mediating the association between downward mobility and mortality.  
195 Depression had the strongest attenuation effect among covariates in Model 3 both among men and  
196 women.

197 Generally, upward SES mobility has a protective effect on health compared to persistently  
198 low SES<sup>28</sup>. In our study, ‘upward’ SSS mobility attenuated mortality risks among women, but not  
199 among men. This pattern is contrary to our original hypothesis that upward mobility would not affect  
200 women’s risk, because even if women graduated from college, this would not be reflected by higher  
201 earnings/income. Many Japanese women of the generation in our sample became housewives  
202 irrespective of their years and quality of education<sup>15</sup>. For men, we did not find either a harmful or  
203 protective effect of ‘upward’ mobility in this study. We can cite two possibilities why upward mobility  
204 is sometimes harmful. Firstly, status incongruity refers to when upward mobility creates stress  
205 because the individual feels like an outsider in her new status/position. Striving for upward mobility

206 and success can also induce stress, also known as the John Henryism phenomenon<sup>29</sup>. Secondly, the  
207 “long arm” of low childhood SES may continue to exert an influence on health even if the individual  
208 manages to escape from lower status, i.e. poor childhood SES persists even if someone achieves high  
209 SES in later life<sup>30</sup>. For women in this study, there was a protective effect of ‘upward’ mobility. In  
210 theory, upward mobility for Japanese women could be even more stressful than in the west, because  
211 it was so non-normative and consequently women in positions of authority and power could become  
212 very isolated. However, in the age-group represented by our cohort, we may assume that the majority  
213 of ‘upward’ mobility more closely reflected their father’s or spouse’s social class rather than their  
214 own<sup>31</sup> so that it may not be indicative of exposure to status incongruity.

215 Our results are also notable for finding that mobility in SSS had a similar impact on mortality  
216 even after adjusting for objective SES. The contribution of SSS to mental and physical health—over  
217 and above the contribution of objective SES—has been previously discussed<sup>5,32,33</sup>. Singh-Manoux et  
218 al. investigated the determinants of SSS in a cohort of middle-aged British civil servants<sup>3</sup>. They  
219 concluded that classic objective SES such as occupational grade, income and education were highly  
220 correlated to SSS, but none of them fully explained the impact of SSS on health. Two possible  
221 explanations for the residual effect of SSS have been proposed<sup>25</sup>: (a) SSS may reflect an individual’s  
222 SES more comprehensively than objective SES by capturing, for example, feelings of financial  
223 security, stocks of wealth, the respect of peers, satisfaction with standard of living, and so on; and (b)  
224 SSS may capture not only the objective social grade, but also people’s sense of their relative position  
225 in society and its impact on health.

226 On the other hand, the possibility of reverse causation between SSS and health, especially  
227 mental health, has also been raised as an alternative explanation<sup>25,34</sup>. That is, depressed individuals  
228 are more likely to make pessimistic evaluations of their past and present circumstances. However, a  
229 strength of our study is that the outcome was objective (mortality), and not self-assessed (such as  
230 depression), thereby avoiding the threat of common method bias.

231 Limitations of this study should be mentioned. First, childhood SSS was assessed via

232 subjective recall. Previous research, however, supports the reliability of retrospective assessments of  
233 childhood SSS using siblings' recall of childhood SSS <sup>35</sup>. Second, our analysis was limited to all-  
234 cause mortality, and cause-specific mortality was not considered. Future studies should examine  
235 cause-specific mortality, such as coronary heart disease, stroke, or cancer, to clarify the effect of  
236 mobility in SSS among older Japanese adults, who are ranked as having the longest life expectancy  
237 in the world <sup>36</sup>. Third, our findings may have been biased by selective survival due to the fact that all  
238 the participants were aged 65-100 years at baseline, i.e. more vulnerable population may have already  
239 died before the start of follow-up. Lastly, the duration of follow-up (three years) was comparatively  
240 short.

241 In conclusion, our study found a possible harmful effect of downward SSS mobility on  
242 mortality among Japanese older men and protective effect of upward SSS mobility among women.  
243 The adverse downward effect was mediated by depression. Additionally, mobility in SSS in later life  
244 may have higher predictive value on mortality related to depression than objective SES. Mobility in  
245 SSS from childhood to older age may be important to be considered when assessing the mortality risk  
246 of older people.

247

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249

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270

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**Table 1 Characteristics of participants**

	Men (n=7,944)	Women (n=8,746)
	N (%)	N (%)
<b>Age (years)</b>		
65–69	2731 (34.4%)	2844 (32.5%)
70–74	2332 (29.4%)	2531 (28.9%)
75–79	1656 (20.8%)	1800 (20.6%)
80 ≤	1225 (15.4%)	1571 (18.0%)
<b>Mobility of SSS*</b>		
Persistently high	2042 (25.7%)	2756 (31.5%)
Downward mobility	1588 (20.0%)	2622 (30.0%)
Upward mobility	2829 (35.6%)	2114 (24.2%)
Persistently low	1485 (18.7%)	1254 (14.3%)
<b>Height<sup>a</sup></b>		
Short	580 (7.3%)	1071 (12.2%)
Middle-short	1331 (16.8%)	2442 (27.9%)
Middle	2498 (31.4%)	2884 (33.0%)
Middle-tall	2066 (26.0%)	1461 (16.7%)
Tall	1092 (13.7%)	389 (4.4%)
Missing	377 (4.7%)	499 (5.7%)
<b>Equivalised income (million yen)</b>		
High (≥ 4.00)	851 (10.7%)	782 (8.9%)
Middle (2.00 - 3.99)	2846 (35.8%)	2569 (29.4%)
Low (< 2.00)	3201 (40.3%)	3506 (40.1%)
Missing	1046 (13.2%)	1889 (21.6%)

**Education (years)**

<b>High (<math>\geq 13</math>)</b>	1654 (20.8%)	971 (11.1%)
<b>Middle (10 - 12)</b>	2528 (31.8%)	2930 (33.5%)
<b>Low (- 9)</b>	3465 (43.6%)	4538 (51.9%)
<b>Other/Missing</b>	297 (3.7%)	307 (3.5%)

**Marital status**

<b>Married</b>	6748 (84.9%)	5056 (57.8%)
<b>Widowed</b>	611 (7.7%)	2846 (32.5%)
<b>Divorced</b>	181 (2.3%)	321 (3.7%)
<b>Unmarried</b>	102 (1.3%)	190 (2.2%)
<b>Others/Missing</b>	302 (3.8%)	333 (3.8%)

**Smoking status**

<b>Non-smoker</b>	1873 (23.6%)	7066 (80.8%)
<b>Smoker/ex-smoker</b>	5516 (69.4%)	674 (7.7%)
<b>Missing</b>	555 (7.0%)	1006 (11.5%)

**Alcohol intake**

<b>Current drinker/Ex-drinker</b>	4778 (60.1%)	1332 (15.2%)
<b>Non-drinker</b>	2698 (34.0%)	6923 (79.2%)
<b>Missing</b>	468 (5.9%)	491 (5.6%)

**Walking time**

<b>&lt; 30 min/day</b>	2354 (29.6%)	2869 (32.8%)
<b><math>\geq 30</math> min/day</b>	5156 (64.9%)	5279 (60.4%)
<b>Missing</b>	434 (5.5%)	598 (6.8%)

**Body weight status (BMI, kg/m<sup>2</sup>)**

<b>Underweight (&lt; 18.5)</b>	424 (5.3%)	699 (8.0%)
<b>Normal (18.5–24.9)</b>	5397 (67.9%)	5749 (65.7%)

<b>Overweight (25.0–29.9)</b>	1612 (20.3%)	1550 (17.7%)
<b>Obesity (<math>\geq 30.0</math>)</b>	128 (1.6%)	228 (2.6%)
<b>Missing</b>	383 (4.8%)	520 (5.9%)

**Depressive symptoms**

<b>No depression (GDS** &lt; 5)</b>	4970 (62.6%)	5130 (58.7%)
<b>Moderate depression (<math>5 \leq \text{GDS} &lt; 10</math>)</b>	1406 (17.7%)	1490 (17.0%)
<b>Depression (<math>10 \leq \text{GDS}</math>)</b>	496 (6.2%)	539 (6.2%)
<b>Missing</b>	1072 (13.5%)	1587 (18.1%)

<sup>a</sup>Height (<155, 155–159.9, 160–164.9, 165–169.9, and  $\geq 170$ cm for men and <145, 145–149.9, 150–154.9, 155–159.9, and  $\geq 160$ cm for women).

\*SSS = Subjective Socioeconomic Status

\*\*GDS = Geriatric Depression Scale

Accept



**Table 2. Hazard ratio of SSS\* mobility, SES\*\* and other covariates among men (N=7,944)**

	Model 1	Model 2	Model 3	Model 4
Persistently high	ref	ref	ref	ref
Downward	<b>1.37 (1.08 - 1.74)</b>	<b>1.35 (1.05 - 1.72)</b>	1.22 (0.95 - 1.56)	1.22 (0.95 - 1.56)
Upward	0.87 (0.69 - 1.09)	0.85 (0.68 - 1.08)	0.83 (0.66 - 1.04)	0.84 (0.66 - 1.05)
Persistently low	0.95 (0.73 - 1.24)	0.91 (0.69 - 1.20)	0.83 (0.63 - 1.10)	0.78 (0.59 - 1.03)
	Model 5	Model 6	Model 7	Model 8
Persistently high	1.05 (0.80 - 1.37)	1.09 (0.83 - 1.44)	1.20 (0.91 - 1.58)	1.28 (0.97 - 1.69)
Downward	1.44 (1.10 - 1.88)	1.47 (1.13 - 1.93)	1.46 (1.12 - 1.91)	1.56 (1.19 - 2.04)
Upward	0.91 (0.71 - 1.18)	0.94 (0.72 - 1.22)	0.99 (0.76 - 1.29)	1.07 (0.82 - 1.39)
Persistently low	ref	ref	ref	ref

Models 1 and 5 were adjusted for age

Models 2 and 6 were additionally adjusted for height, equivalised income, education, marital status

Models 3 and 7 were additionally adjusted for GDS\*\*\*

Models 4 and 8 were additionally adjusted for smoking status, drinking habits, BMI\*\*\*\* and walking times

\* SSS = Subjective Socioeconomic Status

\*\* SES = Socioeconomic Status

\*\*\* GDS = Geriatric Depression Scale

\*\*\*\* BMI = Body Mass Index

**Table 3.** Hazard ratios of SSS\* mobility, SES\*\* and other covariates among women (N=8,746)

	Model 1	Model 2	Model 3	Model 4
Persistently high	ref	ref	ref	ref
Downward	1.27 (0.94 - 1.71)	1.22 (0.90 - 1.65)	1.10 (0.81 - 1.50)	1.08 (0.79 - 1.47)
Upward	0.71 (0.49 - 1.04)	0.69 (0.48 - 1.02)	0.68 (0.47 - 1.00)	0.69 (0.47 - 1.01)
Persistently low	1.33 (0.91 - 1.94)	1.20 (0.81 - 1.76)	1.07 (0.72 - 1.58)	1.01 (0.68 - 1.50)
	Model 5	Model 6	Model 7	Model 8
Persistently high	0.75 (0.52 - 1.10)	0.83 (0.57 - 1.23)	0.94 (0.63 - 1.38)	0.99 (0.67 - 1.47)
Downward	0.95 (0.66 - 1.38)	1.02 (0.70 - 1.48)	1.03 (0.71 - 1.49)	1.06 (0.73 - 1.55)
Upward	<b>0.54 (0.35 - 0.83)</b>	<b>0.58 (0.38 - 0.90)</b>	<b>0.64 (0.41 - 0.99)</b>	0.68 (0.44 - 1.07)
Persistently low	ref	ref	ref	ref

Models 1 and 5 were adjusted for age

Models 2 and 6 were additionally adjusted for height, equivalised income, education, marital status

Models 3 and 7 were additionally adjusted for GDS\*\*\*

Models 4 and 8 were additionally adjusted for smoking status, drinking habits, BMI\*\*\*\* and walking times

\* SSS = Subjective Socioeconomic Status

\*\* SES = Socioeconomic Status

\*\*\* GDS = Geriatric Depression Scale

\*\*\*\* BMI = Body Mass Index